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San Diego, California 92152

Statistics of a Chi-Square Random Variable  
Obtained from Independent Gaussian Samples  
with a Non-Zero Mean and Arbitrary Variance

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# Statistics of a chi-square random variable obtained from independent Gaussian samples with a non-zero mean and arbitrary variance

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## ABSTRACT

The mean and variance of a chi-square random variable are generally given for the case in which the chi-square random variable is derived from a process having a zero mean and unit variance. In this report, the mean and variance of the random variable found by squaring and summing  $N$  samples of an independent Gaussian process with a non-zero mean and arbitrary variance is derived.

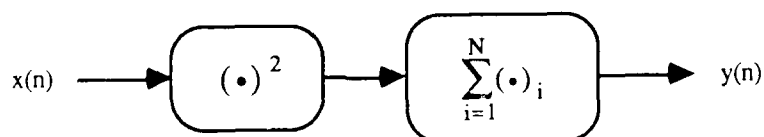
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# Statistics of a chi-square random variable obtained from independent Gaussian samples with a non-zero mean and arbitrary variance

## 1. Introduction

In this report, general expressions are derived for the mean and variance of a random variable that is found by squaring and summing  $N$  independent samples of a Gaussian process  $x_i$ , which has a non-zero mean and an arbitrary variance (i.e.  $\sum_{i=1}^N x_i^2$ ).

A common structure which generates such a random variable is an energy detector (Figure 1). Each output sample is obtained by squaring and summing  $N$  input samples. This is the optimal detector for an unknown signal buried in white Gaussian noise, and is often used as a post-processor for other routines (for example, the output of a beamformer may be run through an energy detector to determine if a signal is present).



**Figure 1.** Energy detector.  
Finds the energy in  $N$  samples of  $x(n)$ . Each output sample is produced by squaring and summing  $N$  input samples.

### *Chi-square random variables*

Let  $z_1, z_2, z_3, \dots$  be normally distributed, independent random variables with zero mean and a variance of one, and define a new random variable

$$\chi_N^2 = z_1^2 + z_2^2 + z_3^2 + \dots + z_N^2 \quad (1)$$

The variable in (1) is called a chi-square random variable with  $N$  degrees of freedom. The density function of  $\chi_N^2$  approaches that of a normally distributed random variable for large  $N$  ( $N > 30$ ), and is non-symmetric for smaller  $N$ . The mean and variance of the chi-square random variable in (1) is given by [1, page 105],

$$\mu_{\chi_N^2} = E\{\chi_N^2\} = N \quad (2)$$

$$\sigma_{\chi_N^2}^2 = \text{var}[\chi_N^2] = 2N \quad (3)$$

Details about the density and distribution functions may be found in [1, Section 4.2.2]. Random variables with a non-zero mean that are squared and summed have a non-central chi-square distribution [3,4].

A time series  $x$  can always be transformed to have a mean of zero and variance equal to one by defining the standardized variable  $z_i$  to be

$$z_i = \frac{x_i - \mu}{\sigma} \quad (4)$$

In some cases, it is desirable to find the statistics of the random variable formed by squaring and summing  $N$  values of  $x$ , which does not necessarily have a zero mean or variance equal to one. In the next section, general expressions for the mean and variance of a random variable that is obtained from squaring and summing independent Gaussian samples with a non-zero mean  $\mu$  and arbitrary variance  $\sigma^2$  are derived.

## 2. Derivation

Given the standardized random variable in (4), a chi-square random variable with  $N$  degrees of freedom is obtained by squaring the  $z_i$  and summing over  $N$  samples,

$$\begin{aligned} \chi_N^2 &= \sum_{i=1}^N z_i^2 \\ &= \frac{1}{\sigma^2} \left[ \sum_{i=1}^N (x_i - \mu)^2 \right] \\ &= \frac{1}{\sigma^2} \left[ \sum_{i=1}^N (x_i^2 - 2\mu x_i + \mu^2) \right] \end{aligned} \quad (5)$$

$$= \frac{1}{\sigma^2} \left[ \sum_{i=1}^N x_i^2 - 2\mu \sum_{i=1}^N x_i + N\mu^2 \right] \quad (6)$$

$$\begin{aligned} E\{\chi_N^2\} &= \frac{1}{\sigma^2} E\left\{\sum_{i=1}^N x_i^2\right\} - \frac{2\mu}{\sigma^2} E\left\{\sum_{i=1}^N x_i\right\} + \frac{N\mu^2}{\sigma^2} \\ &= \frac{1}{\sigma^2} E\left\{\sum_{i=1}^N x_i^2\right\} - \frac{N\mu^2}{\sigma^2} \end{aligned}$$

Solving for the  $N$  squared and summed samples of  $x_i$ , and using Equation (2) gives

$$\begin{aligned} E\left\{\sum_{i=1}^N x_i^2\right\} &= N\sigma^2 + N\mu^2 \\ &= N(\sigma^2 + \mu^2) \end{aligned} \quad (7)$$

By definition, the variance of the chi-square random variable is given by

$$\text{var}[\chi_N^2] = E\{[\chi_N^2 - \mu_{\chi_N^2}][\chi_N^2 - \mu_{\chi_N^2}]\} \quad (8)$$

Using (2) and (6), and multiplying both sides by  $\sigma^4$  gives

$$\sigma^4 \text{var}[\chi_N^2] = E\left\{\left[\sum_{i=1}^N x_i^2 - 2\mu \sum_{i=1}^N x_i + N\mu^2 - N\sigma^2\right]\left[\sum_{j=1}^N x_j^2 - 2\mu \sum_{j=1}^N x_j + N\mu^2 - N\sigma^2\right]\right\}$$

Replacing the expression on the left side with (3) and expanding yields

$$\begin{aligned} 2N\sigma^4 &= E\left\{\left(\sum_{i=1}^N x_i^2\right)^2 - 2\mu \sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j + N\mu^2 \sum_{i=1}^N x_i^2 - N\sigma^2 \sum_{i=1}^N x_i^2 - 2\mu \sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j \right. \\ &\quad + 4\mu^2 \left(\sum_{i=1}^N x_i\right)^2 - 2N\mu^3 \sum_{i=1}^N x_i + 2N\mu\sigma^2 \sum_{i=1}^N x_i + N\mu^2 \sum_{i=1}^N x_i^2 - 2N\mu^3 \sum_{i=1}^N x_i \\ &\quad \left. + N^2\mu^4 - N^2\mu^2\sigma^2 - N\sigma^2 \sum_{i=1}^N x_i^2 + 2N\mu\sigma^2 \sum_{i=1}^N x_i - N^2\mu^2\sigma^2 + N^2\sigma^4\right\} \\ &= E\left\{\left(\sum_{i=1}^N x_i^2\right)^2 - 4\mu \sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j + 2N\mu^2 \sum_{i=1}^N x_i^2 - 2N\sigma^2 \sum_{i=1}^N x_i^2 - 4N\mu^3 \sum_{i=1}^N x_i \right. \\ &\quad \left. + 4N\mu\sigma^2 \sum_{i=1}^N x_i - 2N^2\mu^2\sigma^2 + N^2\mu^4 + N^2\sigma^4 + 4\mu^2 \left(\sum_{i=1}^N x_i\right)^2\right\} \end{aligned}$$

Using Equations (A1) - (A4) from the Appendix in the above equation results in

$$\begin{aligned}
 2N\sigma^4 &= E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} - 4\mu(2N\mu\sigma^2 + N^2\mu\sigma^2 + N^2\mu^3) + 2N^2\mu^2(\sigma^2 + \mu^2) \\
 &\quad - 2N^2\sigma^2(\sigma^2 + \mu^2) - 4N^2\mu^4 + 4N^2\mu^2\sigma^2 - 2N^2\mu^2\sigma^2 + N^2\mu^4 \\
 &\quad + N^2\sigma^4 + 4\mu^2(N\sigma^2 + N^2\mu^2) \\
 &= E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} - 8N\mu^2\sigma^2 - 4N^2\mu^2\sigma^2 - 4N^2\mu^4 + 2N^2\mu^2\sigma^2 + 2N^2\mu^4 \\
 &\quad - 2N^2\sigma^4 - 2N^2\mu^2\sigma^2 - 4N^2\mu^4 + 4N^2\mu^2\sigma^2 + N^2\mu^4 - 2N^2\mu^2\sigma^2 + N^2\sigma^4 \\
 &\quad + 4N\mu^2\sigma^2 + 4N^2\mu^4 \\
 &= E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} - 4N\mu^2\sigma^2 - 2N^2\mu^2\sigma^2 - N^2\mu^4 - N^2\sigma^4
 \end{aligned}$$

Therefore,

$$E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} = 2N\sigma^4 + 4N\mu^2\sigma^2 + 2N^2\mu^2\sigma^2 + N^2\mu^4 + N^2\sigma^4 \quad (9)$$

For a random variable  $y$ ,

$$\text{var}[y] = E\{y^2\} - (E\{y\})^2$$

therefore the variance of  $\sum_{i=1}^N x_i^2$  is given by

$$\text{var} \left[ \sum_{i=1}^N x_i^2 \right] = E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} - \left( E \left\{ \sum_{i=1}^N x_i^2 \right\} \right)^2$$

Using Equations (9) and (7) results in

$$\begin{aligned}
 \text{var} \left[ \sum_{i=1}^N x_i^2 \right] &= E \left\{ \left( \sum_{i=1}^N x_i^2 \right)^2 \right\} - N^2(\sigma^2 + \mu^2)^2 \\
 &= 2N\sigma^4 + 4N\mu^2\sigma^2 + 2N^2\mu^2\sigma^2 + N^2\mu^4 + N^2\sigma^4 - N^2\sigma^4 - 2N^2\mu^2\sigma^2 - N^2\mu^4
 \end{aligned}$$

$$= 2N\sigma^4 + 4N\mu^2\sigma^2 \quad (10)$$

### 3. Summary

If  $x$  is a Gaussian random process and has mean  $\mu$  and variance  $\sigma^2$ , then

$$E\left\{\sum_{i=1}^N x_i^2\right\} = N(\sigma^2 + \mu^2)$$

$$var\left[\sum_{i=1}^N x_i^2\right] = 2N\sigma^4 + 4N\mu^2\sigma^2$$

Note that when  $\mu=0$  and  $\sigma^2=1$  these equations reduce to (2) and (3).



### Appendix: Expected values of various summations

$$1. \quad E \left\{ \sum_{i=1}^N x_i \right\} = N\mu \quad (A1)$$

$$2. \quad E \left\{ \sum_{i=1}^N x_i^2 \right\} = N(\sigma^2 + \mu^2) \quad (A2)$$

This is Equation (7) and was derived in the main text.

$$3. \quad E \left\{ \left( \sum_{i=1}^N x_i \right)^2 \right\} = N\sigma^2 + N^2\mu^2 \quad (A3)$$

#### *Derivation*

Let  $x_i$  be normally distributed random variables and define  $\bar{x} = \sum_{i=1}^N x_i$ , then  $\bar{x}$  is a normal random variable with mean  $N\mu$  and variance  $N\sigma^2$ . The variance of  $\bar{x}$  may be written as

$$\text{var}[\bar{x}] = E\{\bar{x}^2\} - (E\{\bar{x}\})^2$$

so that

$$\begin{aligned} E\{\bar{x}^2\} &= E\left\{ \left( \sum_{i=1}^N x_i \right)^2 \right\} \\ &= \text{var}(\bar{x}) + (E\{\bar{x}\})^2 \\ &= N\sigma^2 + N^2\mu^2 \end{aligned}$$

$$4. \quad E \left\{ \sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j \right\} = 2N\mu\sigma^2 + N^2\mu\sigma^2 + N^2\mu^3 \quad (A4)$$

#### *Derivation*

$$E \left\{ \sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j \right\} = E \left\{ \sum_{i=1}^N \sum_{j=1}^N x_i^2 x_j \right\} = \sum_{i=1}^N \sum_{j=1}^N E \{ x_i^2 x_j \}$$

When  $i = j$  this becomes the single sum  $\sum_{i=1}^N E\{x_i^3\}$ . From [2, page 162],

$$E\{x_i^3\} = 3\mu\sigma^2 + \mu^3$$

so

$$\sum_{i=1}^N E\{x_i^3\} = N(3\mu\sigma^2 + \mu^3) \quad i=j$$

When  $i \neq j$

$$\sum_{i=1}^N \sum_{j=1}^N E\{x_i^2 x_j\} = \sum_{i=1}^N \sum_{j=1}^N E\{x_i^2\} E\{x_j\}$$

since the processes are independent. Using  $E\{x_i\} = \mu$ , and

$$E\{x_i^2\} = \text{var}\{x\} + \mu^2 = \sigma^2 + \mu^2$$

results in

$$\sum_{i=1}^N \sum_{j=1}^N E\{x_i^2\} E\{x_j\} = \sum_{i=1}^N \sum_{j=1}^N \mu(\sigma^2 + \mu^2)$$

These are summed over all  $i$  and  $j$  *except* for the case  $i = j$ , so there are  $N(N-1)$  of them, giving

$$\sum_{i=1}^N \sum_{j=1}^N E\{x_i^2\} E\{x_j\} = N(N-1)\mu(\sigma^2 + \mu^2) \quad i \neq j$$

Adding the cases for  $i = j$  and  $i \neq j$  together gives

$$\begin{aligned} E\left\{\sum_{i=1}^N x_i^2 \sum_{j=1}^N x_j\right\} &= N(3\mu\sigma^2 + \mu^3) + N(N-1)\mu(\sigma^2 + \mu^2) \\ &= 2N\mu\sigma^2 + N^2\mu\sigma^2 + N^2\mu^3 \end{aligned}$$

### References

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